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FORT KNOX, KENTUCKY

REPORT NO. 575

A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION, A

James N. Cronholm, M.S.

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UNITED STATES ARMY MEDICAL RESEARCH AND DEVELOPMENT COMMAND

7 June 1963

FOR ERRATA

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THE FOLLOWING PAGES ARE CHANGES

TO BASIC DOCUMENT

ERRATUM

Report No. 575

A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION, A

James N. Cronholm, M. S.

Page 1, 2nd paragraph, line 1, should read: Inspection of equation (will show that A is a one-to-one func-

Report Submitted 25 March 1963

Author

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	ACCESSION NO. US Army Medical Research Lab, Ft. Knox, Ky. A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION. H J. N. Cronholm Report No. 575, 7 Jun 63, 9 pp & i - 1 table DA Project No. 3A012001B801, Unclassified Report A two-variable generating function is described which yields the sampling probabilities of the Shannon-Wiener information measure. Expension and collection of terms in like powers of the first variable
ties ties	purposes the restriction that the sum of the k category frequencies equal n; collection of terms in like powers of the second variable then produces terms whose coefficients are the required probabilities. The method may be used with either equal or unequal category probabilities for any finite n and k, and thus represents a general solution to the small sample problem. Tables of sampling probabilities are presented.

REPORT NO. 575

A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION, Ĥ

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7 June 1963

Basic Research in Psychological and Social Sciences
DA Project No. 3A012001B801

Report No. 575
DA Project No. 3A012001B801

ABSTRACT

A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION, Ĥ

OBJECT

To describe a method of obtaining the exact sampling probabilities of the maximum likelihood estimate of the Shannon-Wiener information measure, and to present tables of these probabilities.

RESULTS

A two-variable generating function was described which yielded the sampling probabilities of the Shannon-Wiener information measure. Expansion and collection of terms in like powers of the first variable imposed the restriction that the sum of the k category frequencies equal n; collection of terms in like powers of the second variable then produced terms whose coefficients were the required probabilities. Tables of these probabilities were presented for the equiprobable case.

CONCLUSIONS

The method may be used with either equal or unequal category probabilities for any finite n and k, and thus represents a general solution to the small sample problem for this widely used statistic.

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A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION, Ĥ

The outcomes of many experiments are characterized by the distribution of n independent observations among k mutually exclusive and exhaustive categories. In situations of this sort, a number of statistics may be computed as functions of the category frequencies, n₁. A popular, and sometimes useful statistic is the maximum likelihood estimate of the Shannon-Wiener measure of information [2,3]:

$$\hat{H} = -\sum_{i=1}^{k} \frac{n_i}{n} \log_2 \frac{n_i}{n}$$

$$-\log_2 n - \frac{1}{n} \sum_{i=1}^{k} n_i \log_2 n_i \qquad , \qquad (1)$$

where

$$\sum_{i=1}^{k} \mathbf{n}_{i} = \mathbf{n} \tag{2}$$

While approximate tests of the significance of \hat{H} are available for large n [1], knowledge of the exact sampling probabilities is essential for accuracy when n is small, and of considerable theoretical interest for any n. The purpose of this paper is to present a general method for finding exact sampling probabilities of \hat{H} using a two-variable generating function.

Inspection of equation (1) will show that \hat{H} is a single valued function of the sum

$$S_r = \sum_{i=1}^k n_i \log_2 n_i \qquad . \tag{3}$$

Accordingly, the probability of obtaining the rth value of \hat{H} is equal to the probability of obtaining the rth value of the sum S, i.e., $f(\hat{H}_r) = f(S_r)$. The two-variable generating function of equation (4) produces the probabilities of these sums:

$$F(t, u) = \prod_{i=1}^{k} \sum_{n_i=0}^{\infty} \frac{1}{n_i!} p_i^{n_i} u^{n_i} t^{n_i \log_2 n_i}.$$
 (4)

The k pis in this function are probabilities subject to the restriction that

$$\sum_{i=1}^{k} p_{i} = 1 , \quad p_{i} > 0 (i=1, 2, \dots, k) ,$$

and may be regarded as the parameters of a multinomial distribution with k categories from which the n independent observations are drawn.

Each of the k sums in F(t, u) may be identified with a category, while the individual terms in each sum represent possible outcomes in the specified category. The general term in the sum representing the ith category,

 $\frac{1}{n_1!} p_i^{n_1} u^{n_1} t^{n_1 \log_2 n_1}$

indicates that this category contains n_i observations which could have occurred in any one of n_i ! ways with a probability of $p_i^{n_i}$, and that the corresponding ith term in the sum, S_r , will be $n_i \log_2 n_i$.

If the multiplication in (4) is carried out and terms in like powers of u are collected, F(t, u) may be expressed as

$$F(t,u) = \sum_{n=0}^{\infty} g_n(t) \frac{u^n}{n!} .$$

This operation invokes the restriction (2) that the sum of the category frequencies equal the number of observations, n. Multiplication of the coefficients of uⁿ by n! accounts for the possible orderings of all n observations.

Next, if terms in like powers of t are collected in $g_n(t)$, the resulting terms will have the form

$$T = \left\langle \sum_{i=1}^{\left[n_{i}\right]_{i}} \sum_{i=1}^{q_{i}} \frac{1}{\prod_{i=1}^{k} n_{i}!} p_{i}^{n_{i}} \right\rangle t^{s_{i}} , \qquad (5)$$

in which the first sum is over the

$$\mathbf{q} = \frac{\mathbf{k}!}{\prod_{i=1}^{n} \mathbf{k}_i!} , \tag{6}$$

with

$$\sum_{j=0}^{n} \mathbf{k}_{j} = \mathbf{k}_{j}$$

possible orderings of a distinct set of n_i s among the k categories. In q (6), the k_j s are the number of n_i s equal j. The second sum in (5) is over all distinct sets of n_i s, $\begin{bmatrix} n_i \end{bmatrix}_r$, which yield the rth value of S.

This expression (5) can be interpreted readily. The multinomial probability within the summation signs gives the probability of obtaining

a particular set of category frequencies in the indicated categories. Summing over the distinct permutations of the k categories gives the probability of obtaining a particular set of n_i s irrespective of category. Finally, summing these probabilities over all distinct sets of n_i s which yield the sum S_r gives the probability of obtaining S_r . Since the observed set of n_i s is included in these sets, the coefficient of t^{S_r} in (5) gives the required probability.

 $g_n(t)$ is a polynomial in t in which the exponents of t are the possible values of S_r , and the coefficients of t are the corresponding probabilities. Thus, $g_n(t)$ may be written

$$g_n(t) = \sum_{r \in I} f(s_r) t^{s_r}$$
,

=
$$\sum_{r=1}^{r} f(\hat{H}_r) t^{s_r}$$

and the sampling probability of the rth value of \hat{H} is given by (5), with t set equal to 1.

Thus,

$$f(\hat{\mathbf{H}}_{t}) = \sum_{i=1}^{[n_{i}]_{t}} \sum_{i=1}^{q} \frac{\mathbf{n}_{t}}{\prod_{i=1}^{k} \mathbf{n}_{t}!} \prod_{i=1}^{k} \mathbf{p}_{t}^{n_{i}} . \tag{7}$$

A special case of some importance arises when the category probabilities are known or assumed to be equal, i.e., $p_i = 1/k$ (i = 1, 2, ..., k). In this case, (7) reduces to

$$f(\hat{H}_r) = \sum_{\substack{[n_t]_r \\ \frac{1}{10}k_t!}} \frac{k!}{\frac{1}{11}n_t!} \frac{n!}{\frac{1}{11}n_t!} \frac{1}{k^n} .$$

In summary, the sampling probabilities of \hat{H} , given n, k, and the population probabilities, p_i , may be determined by first expanding the two-variable generating function, F(t, u) (4), collecting terms in $u^n/n!$ to find $g_n(t)$, and finally, collecting terms in this function in like powers of t to find $f(\hat{H})$.

For computational purposes, it should be noted that if the infinite series in F(t, u) are truncated at $n_i = c$, the functions $g_n(t)$ will be correct for all values of n up to and including c. Terms in higher powers of u than c should be discarded.

As an example of how the two-variable generating function is used consider the equiprobable case in which k = 4, $p_i = 1/k = 1/4$, and

n = 4. Substituting these values in F(t, u) (4), and truncating at $n_i = 4 = c$ gives

$$F(t,u) = (1 + \frac{1}{4}U + \frac{1}{2!} \frac{1}{4!}U^2 t^{2log_2 2} + \frac{1}{3!} \frac{1}{4!}U^3 t^{3log_2 3} + \frac{1}{4!} \frac{1}{4!}U^4 t^{4log_2 4})^4$$

Expanding and collecting terms in uⁿ/n! up to u⁴/4! yields

$$\begin{split} F(t,u) &= 1 + u + (\frac{12}{16} + \frac{4}{16}t^2) \frac{y^2}{2!} + (\frac{24}{64} + \frac{36}{64}t^2 + \frac{4}{64}t^{3\log_2 3}) \frac{y^3}{3!} \\ &+ (\frac{24}{256} + \frac{144}{256}t^2 + \frac{36}{256}t^4 + \frac{48}{256}t^{3\log_2 3} + \frac{4}{256}t^8) \frac{U^4}{4!} \end{split}$$

in which the polynomial coefficients of $u^n/n!$ are the functions $g_n(t)$:

- $-g_1(t)=t^0$
- $g_2(t) = \frac{3}{4}t^0 + \frac{1}{4}t^2$
- $g_3(t) = \frac{3}{8}t^0 + \frac{9}{16}t^2 + \frac{1}{16}t^{3\log_2 3}$
- $g_4(t) = \frac{3}{32}t^0 + \frac{9}{16}t^2 + \frac{9}{64}t^4 + \frac{3}{16}t^{3\log_2 3} + \frac{1}{64}t^8$

The exponents of t in these expressions are the possible values of S_r for n=1,2,3,4, and k=4. The value of \hat{H} determined by each S_r may be found by substituting S_r (3) in the equation for \hat{H} (1). For $g_4(t)$ the S_r s are: $S_1=.0000$, $S_2=2.0000$, $S_3=4.0000$, $S_4=3\log_2 3=4.7549$, and $S_5=8.0000$. The corresponding values of \hat{H} are: $\hat{H}_1=2.0000$, $\hat{H}_2=1.5000$, $\hat{H}_3=1.0000$, $\hat{H}_4=.8113$, and $\hat{H}_5=.0000$. The sampling probabilities associated with these values of \hat{H} are given by the coefficients of t^{S_r} in $g_4(t)$. Thus, f(2.0000)=3/32, f(1.5000)=9/16, f(1.0000)=9/64, f(.8113)=3/16, and f(.0000)=1/64. The functions $g_1(t)$, $g_2(t)$, and $g_3(t)$ may be used in the same way to identify the sampling probabilities of \hat{H} when n=1,2, and 3.

Table 1 is the result of similar calculations for the special case of equiprobable categories, $2 \le k \le 11$, and $2 \le n \le 12$. The first page of Table 1 and the entries for $9 \le n = k \le 11$ were computed by hand, set aside, and recomputed. As a further check, the expected values of these sampling distributions were compared with the values obtained by Rogers and Green [2]. At this point the probabilities for the entire table became available through computer-expansion of the generating function, and the hand calculations were used to check these results.*

^{*}The author gratefully acknowledges the contribution of A. T. Chen of the University of Louisville Speed Scientific School Computer Laboratory who developed the program used to expand the generating function.

The exact sampling probability of an obtained \hat{H} may be found by entering Table 1 at n, k, and the value of the observed \hat{H} . Tests of the significance of obtained values of \hat{H} under the hypothesis of equal category probabilities may be devised by summing the tabled sampling probabilities over appropriately chosen rejection regions.

While the use of Table 1 is restricted to the case of equal category probabilities, it should be emphasized that the generating function F(t, u) (4) may be employed in the general case of unequal probabilities for any finite n and k. The proposed method therefore represents a general solution to the small sample problem.

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Rogers, M. S. and B. F. Green. The moments of sample information when the alternatives are equally likely. In H. Quastler (Ed.), Information theory in psychology. Glencoe, Ill.: The Free Press, pp 101-107, 1955.

 $TABLE \ 1$ Exact sampling probabilities of \hat{H} for $2 \le k \le 11$, $2 \le n \le 12$, and $p_i = 1/k$.

-	-					Κ.					
N	A	2	3	4	5	6	7	8	9	10	11
2	1.000	.5000 .5000	.6667 .3333	.7500 .2500	.80C0 .20C0	.8333 .1667	.8571 .1428	.8750 .1250	.8889	.9000 .1000	•9091 •0909
3	1.585 .918 .000	•7500 •2500	•2222 •6667 •1111	.3750 .5625 .0625	.4800 .4800 .0400	•5556 •4167 •0278	.6122 .3673 .0204	.6562 .3281 .0156	.6914 .2963 .0123	.7200 .2700 .0100	.7438 .2479 .0083
4	2.000 1.500 1.600 .811	•3750 •5000 •1250	.4444 .2222 .2963 .0370	.0937 .5625 .1406 .1875 .0156	.1920 .5760 .0960 .1280 .0080	.2778 .5556 .0694 .0926 .0046	•3498 •5248 •0525 •0700 •0029	.4102 .4922 .0410 .0547 .0020	.4609 .4609 .0329 .0439 .0014	•5040 •4320 •0270 •0360 •0010	.5409 .4057 .0225 .0301 .0008
5	2.322 1.922 1.522 1.371 .971 .722	.6250 .3125 .0625	.3704 .2469 .2469 .1234	.2344 .3516 .2344 .1172 .0586	.0384 .3840 .2880 .1920 .0640 .0320 .0016	.0926 .4630 .2315 .1543 .0386 .0193 .0008	.1499 .4998 .1874 .1249 .0250 .0125	•2051 •5127 •1538 •1025 •0171 •0085 •0002	.2561 .5121 .1280 .0854 .0122 .0061 .0002	.3024 .5040 .1080 .0720 .0090 .0045 .0001	.3442 .4918 .0922 .0615 .0068 .0034
6	2.585 2.252 1.918 1.792 1.585 1.459 1.252 1.000 .018	.3125 .4687 .1875 .J312	.1234 .4938 .1234 .0823 .1234 .0494	• 2637 •1172 •0879 •3516 •0879 •0293 •0439 •0176 •0010	.1152 .3456 .1536 .0576 .2304 .0576 .0128 .0192 .0076	.0154 .2315 .3472 .1343 .0386 .1243 .0386 .0064 .0096 .0096	.0428 .3213 .3213 .1428 .0268 .1071 .0268 .0036 .0054 .0021	.0769 .3845 .2884 .1282 .0192 .0769 .0192 .0021 .0032 .0015	.1138 .4268 .2561 .1138 .0142 .0569 .0142 .0014 .0020 .0008	.1512 .4536 .2268 .1008 .0108 .0108 .04.2 .0108 .0009 .0013 .0005	1878 4694 2012 0894 0004 0004 0006 0009
7	2.807 2.522 2.236 2.128 1.950 1.842 1.664 1.557 1.449 1.379 1.149 985 .863 .592	•5469 •3281 •1094 •∪156	.2881 .1920 .2881 .0560 .0576 .0192	•1538 •3076 •0513 •1538 •1025 •1538 •0358 •0454 •0451 •0602	.1612 .0537 .1612 .3225 .0537 .0806 .0537 .0806 .0161 .0089 .0054 .0018	.0540 .2701 .0900 .1350 .2701 .0450 .0450 .0300 .0450 .0038 .0022 .008	.0061 .1285 .3213 .1071 .1071 .2142 .0357 .0268 .0178 .0268 .0054 .0018 .0011	.0192 .2019 .3364 .1122 .0841 .1682 .0280 .0168 .0112 .0168 .0034 .0009 .0006	.0379 .2655 .3319 .1106 .0664 .1328 .0221 .0111 .0074 .0111 .0022 .0005 .0003 .0001 .000	.0604 .3175 .3175 .1058 .0529 .1058 .0176 .0075 .0075 .0005 .0005 .0005 .0005 .0005	. Gaba . 3505 . 2901 . 0946 . 0427 . 085 . 0142 . 005 . 0001 . 0000 . 0000 . 0000
8	3.000 2.750 2.500 2.406 2.250 2.156 2.000 1.906 1.811 1.750 1.561 1.5549 1.500 1.406 1.299 1.061 1.000 954 811	.2734 .4375 .2187 .0625	.2561 .1920 .2561 .1536 .0256 .0320 .0512 .0256 .0073	.0384 3.076 11025 .1538 .1025 .0205 .0769 .1025 .0615 .0102 .0051 .0051	.1290 .1720 .U537 .2586 .U860 .1290 .U430 .0172 .0322 .0430 .U014 .U025 .U014 .U014 .U014 .U004	.0900 .0240 .1800 .2401 .0525 .1800 .0600 .0900 .0200 .0120 .0120 .0120 .0120 .0120 .0120 .0120 .0120 .0120 .0120 .0120 .0120	.0245 1836 .0490 .1836 .2448 .0429 .1224 .0102 .0102 .0061 .0076 .0062 .0061 .0010 .0002 .0001	.0024 .0673 .2523 .0673 .1682 .2243 .0386 .0841 .0056 .0042 .0056 .0042 .0056 .0001 .0000 .0001		.0181 .1693 .0175 .0847 .1270 .1693 .0164 .0212 .0020 .0028 .0015 .0020 .0015 .0020 .0016	0910 2172 3227 086 10 6 18 18 0013 0015 0013 0024 0001 0000 0000 0000

TABLE 1 (Continued) $\label{eq:exact sampling probabilities of \hat{H} for $2 \le k \le 11$, $2 \le n \le 12$, and $p_i = 1/k$.$

						K					
N	Ĥ	2	3	4	5	6	7	8	9	10	11
9	3.170								• 0009	.0036	.008
	2.725						•0472	.0108	.0337	.0653	.1016
	2.642						.0105	•1136 •0252	•1770 •0393	•2286	• 2666
	2.503					•0900	.1574	.1892	•1967	.0508 .1905	·0592
	2.419	1				.0900	.1574	. 1892	.1967	1905	.1778
	2.281				.0581	.0765	.0748	.0662	.0566	.0476	0400
	2.197	1			.2322	.2701	.2361	. 1892	.1475	.1143	.0889
	2.113 2.059	1			.0516	.0600	.0524	.0421	.0328	.0254	.0198
	1.975	J		•1154	• 0774 • 07.4	.0900 .0450	•0787	.0631	• 0492	.0381	.0296
	1.891			.2307	1548	.0900	•0262 •0524	.0158 .0315	• 0098 • 0197	•0064	.0042
	1.880	Į			.0077	.0090	.0079	.0063	• 0049	.0127 .0038	.003
	1.837			•1730	.1161	.0675	.0393	.0236	.0148	•0095	• 0063
	1.753	ŀ		•1154	.0774	.0450	.0262	.0158	.0098	.0064	•0042
	1.658	1		•0692	.0464	•0270	•0157	.0095	.0059	.0038	.002
	1.585		.0854	•0256	.0086	.0033	.0014	•0007	.0004	.0002	.0001
	1.447	l	.3841	-1154	.0387	.0150	•0066	.0032	.0016	•0009	.000
	1.435		.1152	.0077 .0346	.0052 .0116	.0030	.0017	•0011	• 0006	.0004	.0003
	1.392	1	.0960	.0288	.0097	.0045 .0038	•0020 •0016	.0009 .0008	. 0005	•0003	•0002
	1.352	ŀ	.1536	.0461	.0155	.0060	.0026	.0013	• 0004 • 0006	•0002 •0004	• 0001
	1.224		.0768	.0231	.0077	.0030	.0013	.0006	.0003	•0002	• 0002 • 0001
	•991	•4922	.0384	.0058	.0013	.0004	.0001	.0001	.0000	.0000	.0000
	•986		• · · 110	·0033	.0011	.0064	.C002	.0001	.0000	.0000	.0000
	•918	• 3281	.0256	.0038	.0009	.0002	.0001	.0000	.0000	.0000	.0000
	•764 •503	.1406	•0110	.0016	. 0664	.0001	•0000	.0000	.0000	.0000	.0000
	•000	.0352 .0039	•0027	.0004	.0001	.0000	.0000	•0000	.0000	•0000	.0000
_		.0039	•0002	•0000	• 0000	•0000	•0000	•0000	• 0000	•0000	.0000
10	3.322								0047	.0004	.0015
	2.922							.0236	.0047	•0163	-0346
	2.846							.0045	.0656 .0125	•1143 •0218	.1616
	2.722						.0562	.1183	. 16.39	1905	.2020
	2.646						.0450	.0946	.1311	.1524	.1616
- 1	2.522					.0563	.0881	.0966	.0929	.0841	.0741
- 1	2.371	l				-1500	.2248	.2366	.2186	.1905	.1616
- 1	2.322	1			.0116	.0250 .0488	.0375	.0394	. 0364	•0318	.0269
1	2.246	ľ			1548	•1500	•0646 •1124	.0651 .0788	.0587	•0505	.0424
- 1	2.171	l			1548	-1500	.1124	.0788	• 0546 • 0546	.0381	.0269
- 1	2.161	l				.0030	0045	•0047	• 0044	.0038	• 0269 • 0032
- 1	2.122	l			.1161	.1125	.0843	0591	.0410	.0286	.0202
- 1	2.046	l			· 0516	.0500	.0375	.0263	.0182	.0127	.0090
	1.971	l		•1442	.0774	.0375	.0187	.0098	.0055	.0032	.0019
ı	1.922	l		•0721	.0310	•0300	.0225	.0158	.0109	.0076	.0054
- 1	1.895	l		•0641	.0387	.0188	.0094	•0049	.0027	.0016	.0010
- 1	1.846	Ī		2884	.0344	.0167	0083	.0044	.0024	.0014	.0008
- 1	1.771			. 2004	.1548 .0026	.0750 .0025	.0375	.0197	.0109	.0064	.0038
- 1	1.761			.0865	.0464	•0225	.0019 .0112	.0013	.0009	.0006	.0004
- 1	1.722	!		.0360	.0194	.0094	.0047	.0025	.0033 .0014	.0019	.0012
	1.685			.0577	.0310	.0150	.0075	.0039	.0022	.0013	.0008
	1.571		.2134	.0769	.0284	.0117	•0053	.0026	.0014	.0008	.0005
	1.522		•1600	.0360	.0097	.0031	.0012	.0005	.0002	.0001	.0001
	1.485		.2561	.0577	.0155	.0050	.0019	.0008	.0004	.0002	.0001
	1.361		.1280	.0144 .0288	.0039 .0077	.0012	0005	.0002	.0001	.0000	.0000
i	1.357			.0027	.0015	.0025	.0009	.0004	• 0002	.0001	.0000
	1.295		.∪854	.0192	•0052	.0007	.0004	.0002 .0003	.0001	.0001	.0000
	1.157		.∪366	.0082	.0022	.0007	.0003	•0001	• 00 01	.0001	.0000
- 1	1.000	.2461	· U128	.0614	. 6002	.0001	.0000	.0000		.0000	.0000
	•971	• 4102	.0213	.0024	.0004	•0001	.0000	•0000	• 0000	•0000	.0000
	•922		.0046	.0010	.0003	0001	.0000	.0000	.0000	•0000	.0000
1	.881	.2344	.0122	.0014	.0002	.0001	.0000	.0000	. 0000	.0000	.0000
	• 722 • 469	.0879 .0195	.0046	.0005	.0001	.0000	.0000	.0000	.0000	.0000	.0000
1	•000	.0020	.0010	.0000	.0000	•0000	.0000	•0000-	.0000	.0000	.0000
			# 000T	• 0000	. 0000	.0000	.0000	.0000	. 0000	.0000	.0000

TABLE 1 (Continued) Exact sampling probabilities of \hat{R} for $2 \le k \le 11$, $2 \le n \le 12$, and $|p_i| = 1/k$.

- [K					
N	Ĥ	2	3	4	5	6	7	8	9	10	11
11	3.459									0020	.0001
- 1	3.278								.0114	.0020 .0359	.0692
	3.096								.0019	.0060	.0115
- 1	2.914							. 0325	.0801	.1257	.1616
	2.845							.0217	.0534	.0838	•107
	2.732						.0442	0529	·1040 ·2003	.2096	· 10e
ľ	2.664						.0883	• 1626 • 0217	.0267	.0279	.026
	2.595		-			.0206	.0442	.0569	.0601	.0576	.052
	2.482					.1375	.1766	. 1626	.1336	.1048	.080
- 1	2.413					.0917	.1178	.1084	.0890	.0698	. 053
- 1	2.404					0400	.0012	.0022 .0813	.0027	.0028 .0524	.002
	2.368				.0426	.0688 .0573	.0883	.0407	.0306	0227	.016
- 1	2.300				.1703	.1375	.0883	.0542	.0334	.0210	.013
- 1	2.222					.0138	.0177	.0163	.0134	.0105	.008
	2.187				.0852	.0688	.0442	.0271	0167	.0105	.006
	2.163				.0378	.0306	.0196	•0120	.0074	.0046	- 003
4	2.118				.1703	.1375 .0∪09	.0883	•0542 •0011	.0009	.0007	.000
1	2.049				.0511	.0412	.0265	.0163	.0100	.0063	.004
	2.005				0142	0115	.0074	.0045	.0028	.0017	.001
	1.981			.0881	.0378	.0153	.0065	.0030	.0015	.0008	.000
	1.972				.0227	.0183	•01 18	.0072	.0044	.0028	.001
	1.936			.1983	.0852	.0344	.0147	•0068	.0033	.0017	.001
	1.868			•1322 •0396	.0681	.0321 .0069	.0157 .0029	.0081 .0014	.0044	.0026	.000
	1.856			.0991	.0426	•0172	•0074	.0034	.0017	.0009	.000
	1.790			1586	.0681	.0275	.0118	.0054	.0027	.0014	.000
	1.686			.0396	.0170	.0069	.0029	.0014	.0007	.0003	.000
	1.677			.0396	.0170	.0069	0029	.0014	.0007	.0003	.000
	1.673				.0008	.0006	.0004	.0002	.0002	.0001	.000
	1.617		.1956	.0264	.0114	.0046	.0020 .0066	.0009 .0002	.0001	.0000	.000
	1.573		.1565	.0264	.0057	.0015	.0005	.0002	.0001	.0000	.000
	1.495		2347	.0396	.0085	.0023	.0007	.0003	.0001	.0000	.000
	1.491			.0113	.0049	.0020	.0008	•0004	•0002	.0001	• 000
	1.435		.1565	.0264	0057	.0015	.0005	.0002 .0001	.0001	.0000	.000
	1.348		.0469	.0079 .0132	.0017	.0008	.0002	.0001	.0000	.0000	.000
	1.322		.0335	.0057	.0012	.0003	.0001	.0000	.0000	.0000	.000
	1.278			.0009	. 0064	.0002	.00Ul	.0000	.0000	.0000	.000
	1.241		.0447	.0076	.0016	.0004	•0001	•0001	.0000	.0000	.000
	1.096		.0168	.0028	.0006	.0002	•0000	.0000	.0000	•0000	.000
	•994 •946	.4512 .3223	.0156	.0013	.0002 .0001	.0000	.0000	•0000	.0000	.0000	• 000
	.866	• 3223	.0019	.0003	.0001	.0000	.0000	.0000	.0000	.0000	. 000
	.845	.1611	.0056	.0005	.0001	.0000	.0000	.0000	.0000	•0000	• 000
	.684	.∪537	.0019	.0662	.0000	.0000	.0000	.0000	.0000	.0000	.000
	.440 .000	.0107	.0004	.0000	.0000	.0000	.0000	.0000	.0000	.0000	.000
		.0010	.0000								
12	3.418									•0054	.000
	3.252									.0008	.002
	3.085								.0178	.0503	.088
	3.022							0205	.0102	.0287	.050
	2.918							.0305 .0488	.0674 .1068	.0961 .1509	.176
	2.855							.0054	.0119	0168	.019
	2.792		-				.0227	.0447	.0579	.0629	.06
	2.689						.1009	.1626	.1781	.1676	. 146
	2.626						.0505	.0813	.0890	.0838	.07
	2.617							.0005	.0010	•0014	.00
	2.585					.0034 .0688	.0416 .0858	.0640	.0690 .0623	.0644 .0482	.036
	2 • 522					.1375	.1514	.1220	.0890	.0629	.044
	2.459						.0061	.0098	.0107	.0101	•00
	2.451					.0688	.0757	.0610	.0445	.0314	.022

TABLE 1 (Continued) Exact sampling probabilities of \hat{R} for $2 \le k \le 11$, $2 \le n \le 12$, and $p_i = 1/k$.

N	Ĥ	2	3	4	5	6	7	8	9	10	1
12	2.396					.0204	•0224	.0181	•0132	•0093	•006
						.0917	.1009	.0813			
	2.292				.0681	.0458	•0256	.0141			.002
		1				.0275	.0303	.0244	.0178		
	2 . 252				.0255	.0229	.0158	.0102	. 0065	•0042	•000
	2.230	1			.0908	.0611	.0336	.0181	•0099	•0056	.002
- 1		1				.0092	.0101	.0081	.0059	•0042	• 003
- 1	2 • 189	1			. 2044	.1375	.0757	.0407	.0223		• 002
	2 • 126				.0681	.0504	.0303	.0176	.0104	•0126	.007
- 1	2 • 117				.0409	.0275	•0151	.0081	.0044	•0063	• 003
- 1	2.085	1			.0511	.0344	.0189	.0102		•0025	• 001
- 1	2.054				.0817	.0550	.0303	.0163	•0056	•0031	• 00 1
- 1	2.000			.0220	.0076	.0025	.0009		.0089	.0050	.002
	1.959			.1983	.0886	.0367	0160	• 0004	•0002	• 0001	• 000
	1.951				.0136	•00 92		•0074	.0037	.0020	.001
	1.947					•0003	• 0050	.0027	.0015	.0008	• 000
	1.918			.0744	.0255	.0086	.0003	.0002	.0002	.0001	.000
	1.896				.0091	.0061	• 0032	.0013	• 0006	.0003	.000
- 1	1.888	ľ		•1190	0409		• 0034	.0018	.0010	.0006	.000
- [1.855	l .		.0991	.0341	.0138	• 0050	.0020	• 0009	.0004	.000
J	1.825			.0793		.0115	• 0042	.0017	.0007	.0003	.000
	1.792			.0198	.0272	.00 92	.0034	•0014	.0006	.0003	.000
	1.784				.0068	.0023	.0008	.0003	• 0001	.0001	.000
	1.781			•1190	•0409	.0138	• 0050	.0020	.0009	.0004	.000
- 1	1.730			0700	.0039	.0026	• 0014	.0008	. 0004	.0002	• 000
- 1	1.650			•0793	.0272	.0092	• 0034	.0014	.0006	•0003	.000
- [1.626			•0119	.0041	.0014	• 0005	.0002	.0001	.0000	• 000
- 1	1.614			•0198	.0068	·0U23	.0008	.0003	. 6001	.0001	.000
- 1	1.585			.0170	.0058	.0U2C	.0007	.0003	.0001	• 0C01	
	1.554		.0652	.0083	.0017	.0005	.0002	.0001	. 0000		• 000
- 1			.3130	0396	· U068	.0015	.0004	.0001	.0000	•0000	.000
- 1	1.551			•0113	• 0039	.0C13	. 0005	.0002	.0001	•0000	. 000
- 1	1.500		.1043	.0132	.0023	.0005	.0001	.0000		•0500	.0000
- 1	1.483		. 6939	.0119	.0020	0004	.0001	.0000	.0000	•0000	.0000
- 1	1.459		.1565	.0198	.0034	.0008	• 0002		.0000	•0000	. 0000
	1 • 418			.0042	.0014	.0005		.0001	• 0000	•0000	.0000
1	1.384		. 0894	.0113	.0019	.0004	• 0 0 0 2	.0001	• 0000	•0000	.0000
	1.325		.0626	.0079	.0014	.0003	• 0001	•0000	• 0000	.0000	• 0000
_ [1.281		. 447	.0057	.0010	.0002	.0001	.0000	• 0000	.0000	.0000
- 1	1.252		.0168	.0021	.0004		.0001	.0000	• 0000	.0000	.0000
- 1	1.208			.0003	0001	.0001	•0000	•0000	• 0000	.0000	.0000
- 1	1.189		.0224	.0028	.0005	•0000	•0000	.0000	• 0000	.0000	. 0000
	1.041		.0074	.0009		.0001	•0000	•0000	• 0000	.0000	.0000
- 1	1.000	.2256	.0052		.0002	.0000	• 0000	.0000	.0000	.0000	.0000
	• 980	. 3867	.0089	0003	•0000	.0000	• 0000	.0000	.0000	.0000	.0000
	918	. 2417		0006	.0001	.0000	.0000	.0000	.0000	.0000	.0000
	817	0241/	.0056	.0004	.0000	.0000	.0000	.0000	• 0000	.0000	• 0000
	811	107	.0007	.0001	.0000	.0000	.0000	.0000	.0000	•0000	• 0000
	650	.1074	•0025	.0002	.0000	.0000	.0000	.0000	.0000	.0000	• 0000
1	414	.0322	.0007	.0000	.0000	.0000	.0000	.0000	• 0000	•0000	
		.0058	• · · 00 1	.0000	.0000	.0000	.0000	•0000	• 0000		• 0000
	.000	.0005	.0000	.0000	.0000	.0000	.0000	.0000	• 0000	.0000	.0000

x, Ky. AMPLING I. Statistics EASURE 2. Non-parametric 3. Information fied Report described which yields the Whener information measure. Ex- ke powers of the first variable of the k category frequencies powers of the second variable s are the required probabilities. ual or unequal category probation thus represents a general so- thus represents a general so- Tables of sampling probabilities	w, Ky. AMPLING I. Statistics 2. Non-parametric 3. Information fied Report described which yields the -Wiener information measure. Ex- ke powers of the first variable of the k category frequencies s are the required probabilities. full or unequal category proba- thus represents a general so- thus represents a general so- thus represents a general so-
US Army Medical Research Lab, Ft. Knox, Ky. US GENERAL METHOD OF OBTAINING EXACT SAMPLING A GENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION, H J. N. Cronholm Report No. 575, 7 Jun 63, 9 pp & i - 1 table A two-var able generating function is described which yields the sampling probabilities of the Shannon-Wiener information measure. Exsampling probabilities of the Shannon-Wiener information measure. Expansion and collection of terms in like powers of the first variable imposes the restriction that the sum of the k category frequencies imposes the restriction that the sum of the required probabilities then produces terms whose coefficients are the required probabilities for any finite n and k, and thus represents a general solution to the small sample problem. Tables of sampling probabilities are presented.	Accessing Medical Research Lab, Ft. Knower Meral Method of Obtaining Exact Sabilities of THE SHANNON-WIENER NFORMATION, H J. N. Cronholm Fird No. 575, 7 Jun 63, 9 pp. 6 into a containing function is on and collection of terms in like in the method may be used with either equication to the small sample problem.
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UNCLASSIFIED US Army Medical Research Lab, Ft. Knox, Ky. UCENERAL METHOD OF OBTAINING EXACT SAMPLING PROBABILITIES OF THE SHANNON-WIENER MEASURE OF INFORMATION. H J. W. Cronholm Report No. 575, 7 Jun 63, 9 pc & i l table DA Project No. 340120018801, Unclassified Report A two-variable generating function is described which yields the sampling probabilities of the Shannon-Wiener information measure. Exsampling probabilities of the Shannon-Wiener information measure. Exsampling probabilities of the Shannon-Wiener information measure imposes the restriction that the sum of the k category frequencies the produces terms whose coefficients are the required probabilities for any finite n and k, and thus represents a general solution to the small sample problem. Tables of sampling probabilities are presented.	Accession No. Imy Medical Research Lab, Ft. Knox, Ky. NERAL METHOD OF OBTAINING EXACT SAMPLING ABILITIES OF THE SHANNON-WIENER MEASURE NFORMATION, H J. N. Cronholm It No. 575, 7 Jun 63, 9 pp. d 1 table! reject No. 3A012001B801, Unclassified Reprovariable generating function is describe ling probabilities of the Shannon-Wiener ses the restriction of terms in like powers produces terms whose coefficients are the method may be used with either equal or un ties for any finite n and k, and thus reprovers on to the small sample problem. Tables of
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